

Trade Disruptions and the Organization of Supply Chains

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Sebastian Heise Federal Reserve Bank of New York

Justin R. Pierce Board of Governors of the Federal Reserve System

Georg Schaur University of Tennessee, Knoxville

Peter K. Schott Yale School of Management & NBER & CEPR

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Motivation

- Global value chains have \uparrow world trade and rich-poor country convergence
- Recent events put these supply chains at risk
- Understanding how firms respond to potential disruption is important for many reasons
- We highlight a novel aspect of importer-supplier interactions: choice of procurement system
- Specifically, buyers' choice of two different systems for assuring high-quality inputs
 - Firms buy on the spot market, perform inspections to verify quality
 - Firms form long-term relationships, pay incentive premia to deter cheating

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- We highlight a novel aspect of importer-supplier interactions: choice of procurement system
- Specifically, buyers' choice of two different systems for assuring high-quality inputs
 - Firms buy on the spot market, perform inspections to verify quality – “A” system
 - Firms form long-term relationships, pay incentive premia to deter cheating – “J” system

What Do We Do?

- Theory
 - Start with Taylor & Wiggins (1997)'s model of domestic procurement
 - TW show that *A* and *J* procurement are optimal solutions to a quality control problem
 - We extend to TW to international procurement; add probability of trade disruption (e.g., trade war)
- Data
 - Develop indicator of *A* vs *J* procurement using transaction-level trade data
 - Use this indicator to examine whether buyers and transactions behave as predicted
 - Analyze US importers' reactions to sizeable reduction in tariff uncertainty with China in 2000
- Quantitative simulations
 - Embed our model in Eaton & Kortum (2002) to analyze GE implications
 - What to expect if the probability of trade disruption rises for all US trade partners?

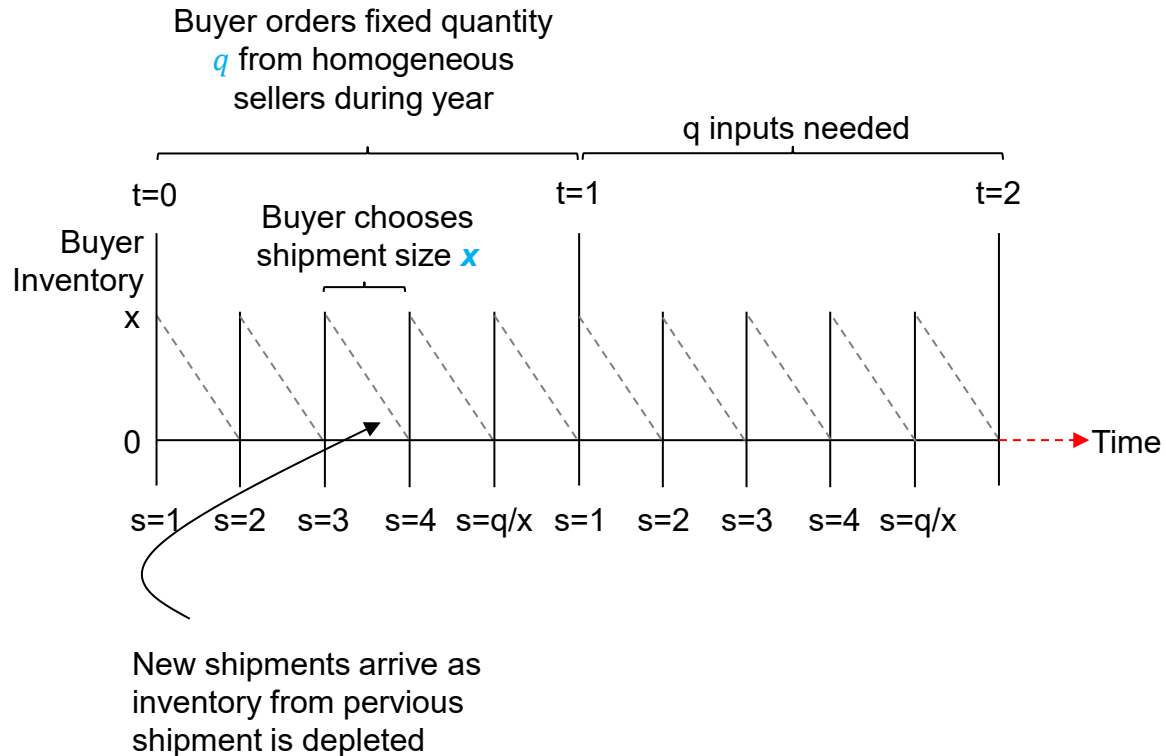
Agenda

- Introduction
- **Model overview**
- Data
- Quantitative simulations

Model Overview

- We start with Taylor and Wiggins' (1997) model of *domestic* procurement
 - Buyer orders inputs from supplier without observing quality
 - Buyer chooses optimal order pattern, payments, and inspections
- TW show that the optimal contract is one of two strategies:
 - *A* (“American” in TW)
 - Buyer chooses lowest cost bidder in spot market
 - Pays fixed inspection cost with some probability (we simplify to 1) to deter cheating
 - Sellers have no bargaining power; thus receive marginal cost
 - *J* (“Japanese” in TW)
 - No inspection; buyer deters cheating via payment of incentive premium in long-term contract
- Our modeling contribution
 - Generalize to int’l procurement, where disruption (e.g., trade war) may break relationships
 - Embed the model into GE (last part of presentation)

Buyer (US Importer) Problem



- Buyer chooses x to minimize the present discounted value of each system
- Cost per order c_i under each system $i \in \{A, J\}$ is

$$c_A(x, q) = \bar{\theta}x + f + \text{Fixed Inspection Cost}$$

Cost per unit of high quality

Fixed order cost

$$c_J(x, q) = \bar{\theta}x + f + \text{Incentive Premium}$$

Per order payment needed to deter seller from cheating; rises with probability of trade disruption

- $\{A, J\}$ choice is trade-off between inspection cost and incentive premium
- Buyer chooses system with lower average costs

Model Implications: Distinguishing Systems

- Order size: $x_A^* > x_J^*$
 - Fixed inspection cost under A implies less frequent orders, so higher order size x
 - Incentive provision under J implies more frequent ordering
- Price: $p_A^* < p_J^*$
 - Incentive premium under J raises the price above marginal costs
- These results also present in TW's domestic procurement
 - Our paper provides the first empirical examination of these predictions

Model Implications: Trade Policy

- Cost of J procurement depends on probability of trade disruption
 - Sellers demand higher incentive not to cheat on quality if disruption is more likely
- Cost of A procurement is independent of trade policy
- Lower probability of trade wars \rightarrow lower incentive premium $\rightarrow J$ becomes less costly
- There exists a unique probability above which A is less costly and below which J less costly
- If the probability of a trade war falls sufficiently, optimal system switches from $A \rightarrow J$
 - Order size falls, since $x_A^* > x_J^*$
 - Price rises, since $p_A^* < p_J^*$

We will see this in US-China trade shortly

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- Introduction
- Model
- Data
 - Data sources
 - Classify importers as A vs J according to number of suppliers
 - Is inferred A vs J buyer procurement consistent with model?
 - Procurement switching after reduction in US-China uncertainty?
- Quantitative simulations

Data Sources

- Transaction-level US import data from 1992 to 2016
 - Observe buyer, seller, value, quantity, departure, arrival
 - Focus on arm's-length transactions
 - Focus on shipments within narrow *mhc* “buyer” quadruples to isolate variation in price, quantity



- Longitudinal Business Database (LBD) and Economic Censuses (ECs)
 - Track various attributes of firms over time, e.g., inventories

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Classifying Importers as *A* vs *J* Using the Model

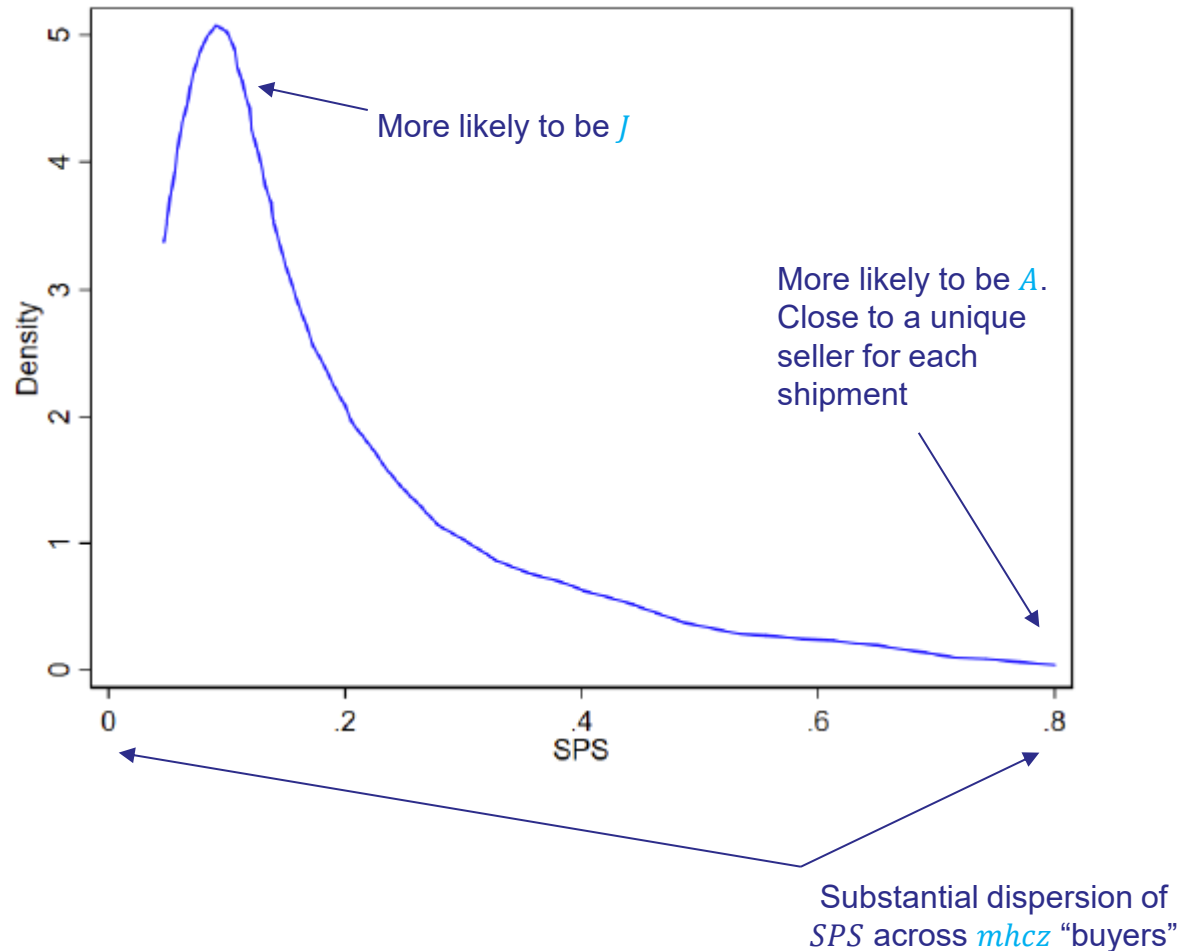
- Model
 - Under *A*, buyers receive shipments from many sellers
 - Under *J*, buyers receive shipments from a single seller
- Data
 - For each *mhcZ* “buyer”, compute sellers per shipment (*SPS*) over some period

$$SPS_{mhcZ}^{beg-end} = \frac{\text{Number of Suppliers}_{mhcZ}^{beg-end}}{\text{Number of Transactions}_{mhcZ}^{beg-end}}$$

- Bounded (0,1]
- $SPS_{mhcZ} = 1$ means different seller for each shipment
- $SPS_{mhcZ} \rightarrow 0$ means many shipments from a given seller
- Higher SPS_{mhcZ} , more likely to be *A* procurement

Classifying Importers as *A* vs *J* Using the Model

Distribution of $SPS_{mhc}^{1996-2016}$



- Average SPS across countries or products to ask:
 - Which countries imports look most *J*?
 - Which products look most *J*?

Mean $SPS_{mhcz}^{beg-end}$ Across Countries and Sectors

1992-2000 vs 2002-2007

Average $SPS_{mhcz}^{beg-end}$ Across c

Country	1995-2000	2002-2007
Mexico	0.095	0.068
Japan	0.107	0.123
Taiwan	0.132	0.114
Canada	0.141	0.120
United Kingdom	0.146	0.225
South Korea	0.156	0.135
France	0.177	0.158
<i>Rest of the World</i>	<i>0.180</i>	<i>0.156</i>
Germany	0.184	0.163
China	0.185	0.147
Brazil	0.190	0.151

Average $SPS_{mhcz}^{beg-end}$ Across h

Product code (HS chapter)	1995-2000	2002-2007
Transportation (86-89)	0.107	0.081
Machinery (84-85)	0.130	0.133
Plastics (39-40)	0.130	0.096
Optical products (90-92)	0.137	0.127
Footwear (64-67)	0.142	0.117
<i>Other products (93-99)</i>	<i>0.151</i>	<i>0.124</i>
Metals (72-83)	0.154	0.128
Food (16-24)	0.155	0.120
Chemicals (28-38)	0.156	0.121
Stones & Jewelry (68-71)	0.159	0.141
Animal products & vegetables (01-15)	0.166	0.132
Minerals (25-27)	0.182	0.203
Leather and wood products (41-49)	0.188	0.153
Textiles (50-63)	0.224	0.177

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Is Inferred *A* vs *J* Buyer Procurement Consistent with Model?

- Model
 - *A* buyers receive larger shipments, less frequently, at lower price
- Data
 - Use SPS_{mhcZ} to compare *A* and *J* shipment attributes via OLS

$$\ln(\bar{Y}_{mhcZ}) = \beta_1 \ln(SPS_{mhcZ}) + \beta_2 \ln(QPW_{mhcZ}) + \beta_3 beg_{mhcZ} + \beta_4 end_{mhcZ} + \lambda_{hcz} + \epsilon_{mhcZ}$$

Shipment size, frequency or unit value

Sellers per shipment (Higher *SPS*, more likely to be *A*)

Quantity per week (hold *q* fixed as in model)

Controls for first and last week of *mhcZ*, and product bin

Procurement Varies with Inferred A,J System as Predicted

	(1) $\ln(QPS_{mhcZ})$	(2) $\ln(WBS_{mhcZ})$	(3) $\ln(UV_{mhcZ})$
$\ln(SPS_{mhcZ})$	0.418*** (0.017)	0.452*** (0.017)	-0.123*** (0.021)
$\ln(QPW_{mhcZ})$	0.701*** (0.014)	-0.308*** (0.014)	-0.287*** (0.020)
Obs	2,966,000	2,966,000	2,966,000
R-squared	0.947	0.674	0.845
Fixed effects	hcz	hcz	hcz
Controls	beg, end	beg, end	beg, end

Note: Standard errors two-way clustered by county and HS chapter.

- Higher buyer SPS_{mhcZ} associated with
 - Bigger shipments (QPS_{mhcZ})
 - More time between shipments (WBS_{mhcZ})
 - Lower unit value (UV_{mhcZ})
- First empirical support for Taylor Wiggins (1997)
- In other results, we find that higher buyer SPS_{mhcZ} also associated with higher inventories

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US-China Sourcing After Substantial ↘ in Tariff Uncertainty

- US grants China “permanent” access to MFN tariffs in 2000
- Exposure to this decline in tariff uncertainty varied by product
- Triple DID empirical strategy
- How did buyers’ shipment size, frequency and unit value respond
 - After versus before PNTR?
 - For US buyer relationships with China vs other countries?
 - In products more vs less exposed to PNTR?

$$\ln(\bar{Y}_{mxhcz}) = \beta_1 Post_t * China_c * Exposure_h + \beta_2 X_{mxhcz} + \beta_3 \ln(QPW_{mxhcz}) + \lambda_{mxhcz} + \lambda_t + \epsilon_{mxhcz}$$

Diagram illustrating the components of the regression model:

- $\ln(\bar{Y}_{mxhcz})$: Shipment size, frequency or unit value
- $\beta_1 Post_t * China_c * Exposure_h$: Triple interaction of Post and China dummies with exposure to PNTR
- $\beta_2 X_{mxhcz}$: Interactions needed to identify β_1
- $\beta_3 \ln(QPW_{mxhcz})$: Quantity per week (hold q fixed as in model)
- λ_{mxhcz} : Fixed effects
- λ_t : Time fixed effects
- ϵ_{mxhcz} : Error term

US-China Sourcing After ↓ in Tariff Uncertainty

	(1) $\ln(QPS_{mxhczt})$	(2) $\ln(WBS_{mxhczt})$	(3) $\ln(UV_{mxhczt})$
$Post_t \times China_c \times Exposure_h$	-0.197*** (0.009)	-0.168*** (0.009)	0.092*** (0.023)
$\ln(QPW_{mxhczt})$	0.368*** (0.009)	-0.632*** (0.008)	-0.124*** (0.013)
Obs	439,000	439,000	439,000
R-squared	0.98	0.89	0.99
Fixed effects	mxhcz, t	mxhcz, t	mxhcz, t
Controls	Y	Y	Y

- After PNTR, more exposed $mxhcz$ “buyer-seller” quintuples exhibit relative
 - ↓ quantity per shipment (QPS_{mxhcz})
 - ↓ weeks between shipments (WBS_{mxhcz})
 - ↑ unit value (UV_{mxhcz})
- These results suggest a transition from $A \rightarrow J$
- Consistent with ↓ in probability of disruption
- Paper contains further evidence

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Quantitative Framework

- We embed our model into a standard EK (2002) model of comparative advantage
 - Quantity traded q is now endogenous
- Challenge: average (and marginal) costs are \downarrow sloping in q due to fixed cost of shipments
 - Standard competitive markets as in EK do not apply
 - So, use Baumol et al. (1982) **contestable markets**: Bertrand competition for markets with IRS
- New sources of heterogeneity
 - Productivity (Also in EK)
 - Bilateral trade peace probabilities (**New**)
 - Country quality inspection costs (**New**)
- Comparative advantage now depends on productivity and policy uncertainty

Quantitative Exercise

- Estimate model with 3 countries for 1992-2016
- US, China, rest of world (ROW)
- Parameterization using mix of calibration and moment matching

Baseline Estimation Results

	Baseline Equilibrium
(1) Share of consumption from China (%)	6.6%
(2) - of which, "Japanese"	9.5%
(3) Share of consumption from ROW (%)	27.6%
(4) - of which, "Japanese"	52.1%
(5) Share of consumption from U.S. (%)	65.8%
(6) Avg. inspection costs	0.4%
(7) Avg. fixed costs (imports)	4.5%
(8) Manufacturing price index	1.000
(9) Utility	1.000

- 9.5 percent of imports from China are *J* vs 52.1 percent for ROW
- ROW share larger due to lower trade war probability and lower fixed costs

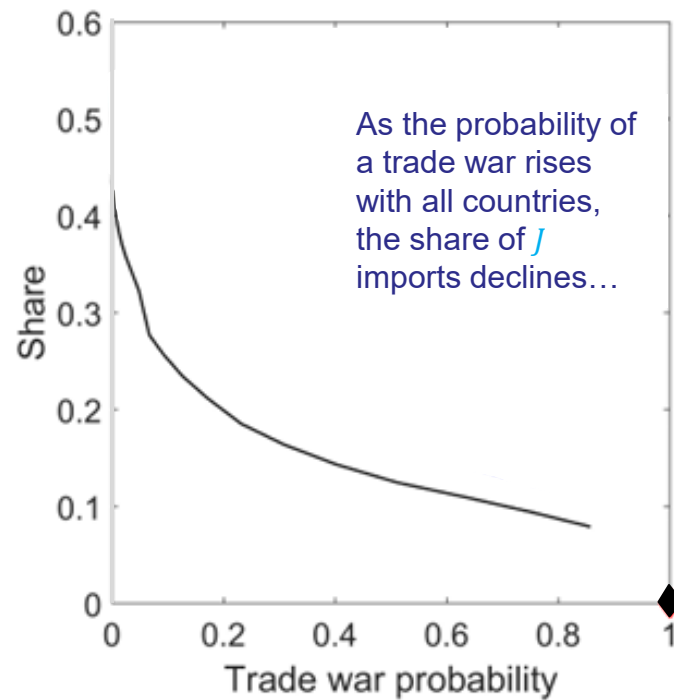
Three Counterfactuals We Examine

- Symmetric increase in the probability of a trade war between the US and all trading partners
- Adoption of trade facilitation policies
- Increased automation

Three Counterfactuals We Examine

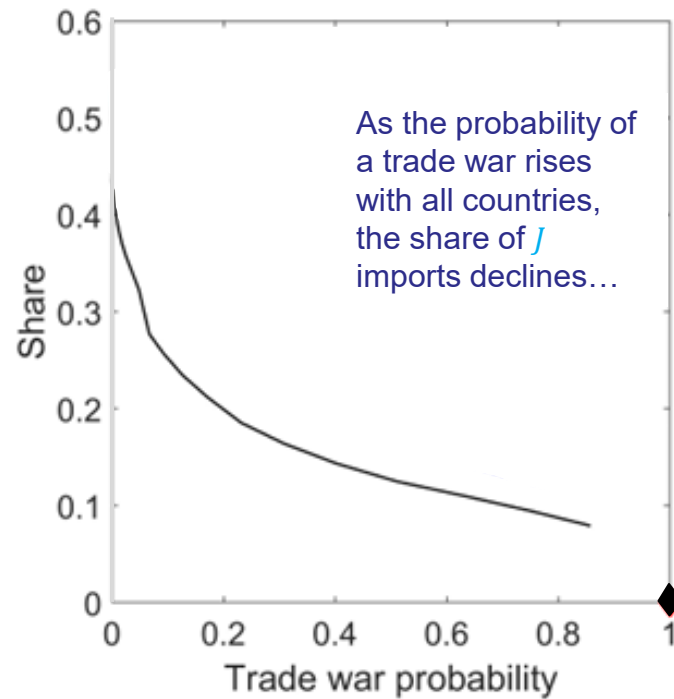
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What if Probability of Trade War Rises Symmetrically For All US Partners?

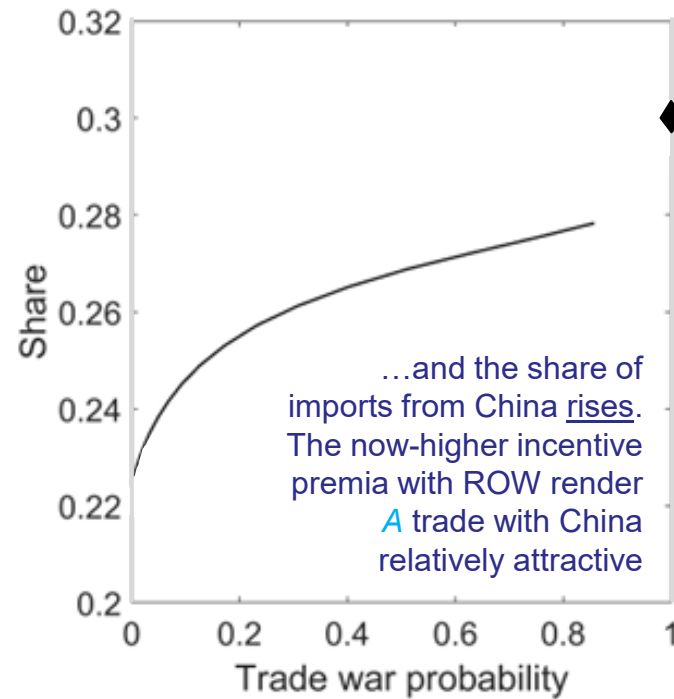


Share of J imports

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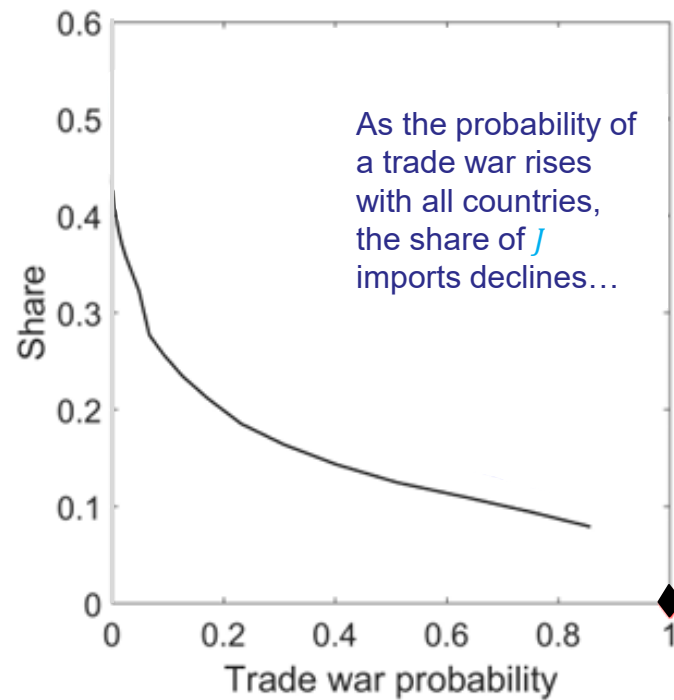


Share of *J* imports

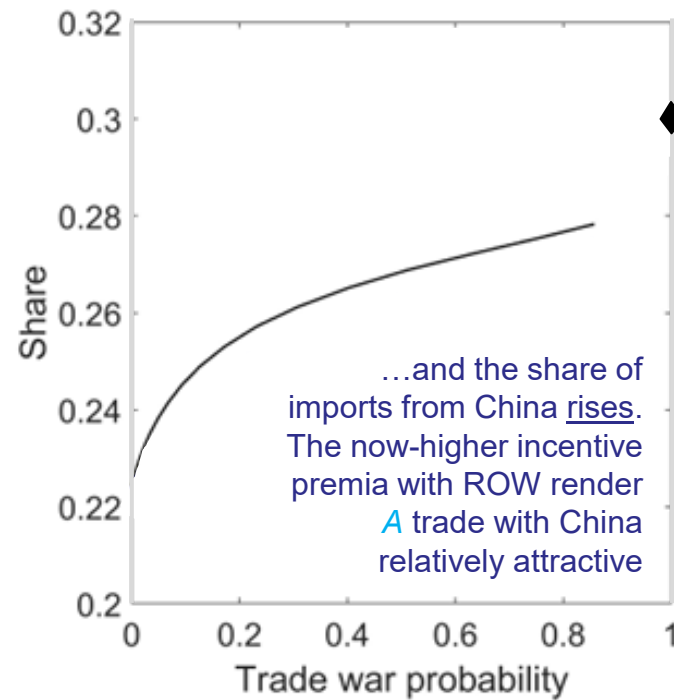


Share of imports from China

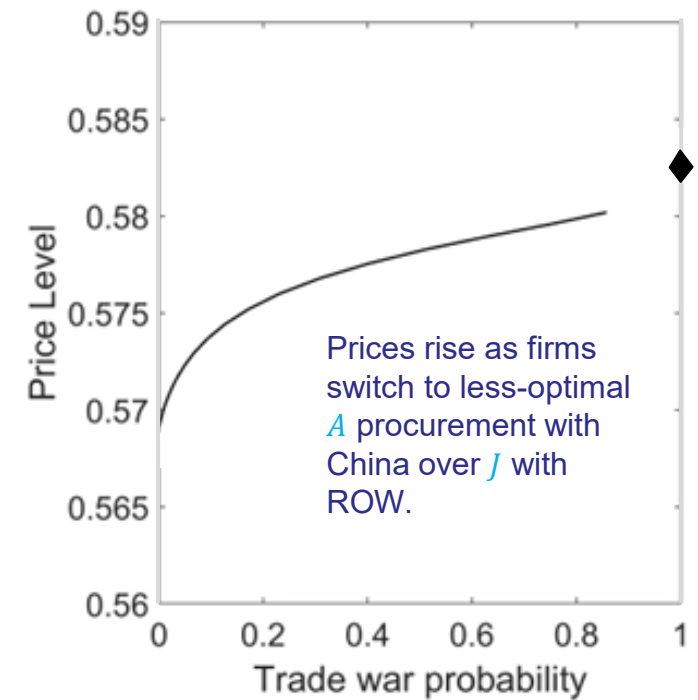
What if Probability of Trade War Rises Symmetrically For All US Partners?



Share of *J* imports



Share of imports from China



Price Level

Conclusion

- We develop a model linking procurement patterns to trade policy uncertainty
- As the probability of trade peace declines, importer-exporter pairs are less likely to form longer-term, “just-in-time” relationships
- Quantitative simulations reveal broad increases in uncertainty can lead to unexpected outcomes

Thanks!